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* Def.: Let X_1, \dots, X_n be n RVs on $(\mathcal{S}, \mathcal{F}, P)$ for some finite $n \geq 1$

$$\underline{X} = [X_1, \dots, X_n]^T \leftarrow \text{Rand. Vec. on } (\mathcal{S}, \mathcal{F}, P)$$

↑
(Random Column Vector)

$$\underline{X}(\omega_1) = \begin{bmatrix} X_1(\omega_1) \\ \vdots \\ X_n(\omega_1) \end{bmatrix}$$
$$\underline{X}(\omega_2) = \begin{bmatrix} X_1(\omega_2) \\ \vdots \\ X_n(\omega_2) \end{bmatrix}$$

* Most definitions & properties for 2 RVs extend naturally to the case of n RVs

* The n^{th} order density fn of a Rand. Vec.:

$$f_{\underline{X}}(\underline{x}) = \frac{\partial^n}{\partial x_1 \dots \partial x_n} F_{\underline{X}}(\underline{x}), \quad \forall \underline{x} = [x_1, \dots, x_n]^T \in \mathbb{R}^n$$

* For any $D \in \mathcal{B}(\mathbb{R}^n)$, $P(\underline{X} \in D) = \int_D f_{\underline{X}}(\underline{x}) d\underline{x}$

(where $\mathcal{B}(\mathbb{R}^n)$ is the sigma field generated by the set of n -dim hyperplanes in \mathbb{R}^n)

* Components of \underline{X} are STATISTICALLY INDEPENDENT if:

$$f_{\underline{X}}(\underline{x}) = \prod_{i=1}^n f_{X_i}(x_i), \quad \forall \underline{x} \in \mathbb{R}^n$$

(for all vectors \underline{x} in \mathbb{R}^n)

Random Vectors: Moments

* Often characterize a random vector via 1st & 2nd order moments

* Correlation btwn X_j & X_k is denoted R_{jk}

$$\Rightarrow R_{jk} \equiv E[X_j X_k], (j, k) = 1, \dots, n$$

$$\text{if } j=k, R_{jj} = R_{kk} = E[X_j^2] = E[X_k^2]$$

* Covariance of X_j & X_k is denoted C_{jk}

$$\Rightarrow C_{jk} \equiv E[(X_j - \bar{X}_j)(X_k - \bar{X}_k)]$$

$$\text{if } j=k, C_{jj} = C_{kk} = \sigma_{X_j}^2 = \sigma_{X_k}^2$$

* Def.: Correlation Matrix for \underline{X} is denoted $R_{\underline{X}}$

$$\textcircled{1} \Rightarrow R_{\underline{X}} = \begin{bmatrix} R_{11} & \dots & R_{1n} \\ \vdots & & \vdots \\ R_{n1} & \dots & R_{nn} \end{bmatrix}$$

← (Diag. contains mean-squared values of X_1 to X_n , off-diagonal terms are the correlations of 2 RVs)

**3 MOMENTS USED
2 CHARACTERIZE
A RANDOM VECTOR**

* Def.: Covariance Matrix for \underline{X} is denoted $C_{\underline{X}}$

$$\textcircled{2} \Rightarrow C_{\underline{X}} = \begin{bmatrix} C_{11} & \dots & C_{1n} \\ \vdots & & \vdots \\ C_{n1} & \dots & C_{nn} \end{bmatrix}$$

← (Diag. terms are the variances of X_1 to X_n , off-diagonal terms are the co-variances)

* Def.: Mean Vector of \underline{X} is denoted $\mu_{\underline{X}}$

$$\textcircled{3} \Rightarrow \mu_{\underline{X}} = [E[X_1], \dots, E[X_n]]^T$$

$$R_{\underline{X}} = E[\underline{X} \underline{X}^T] \leftarrow \begin{matrix} (n \times 1)(1 \times n) \\ (n \times n) \end{matrix}$$

$$C_{\underline{X}} = E[(\underline{X} - \mu_{\underline{X}})(\underline{X} - \mu_{\underline{X}})^T] \leftarrow \begin{matrix} (n \times 1) & (1 \times n) \\ & = (n \times n) \end{matrix}$$

Random Vectors: Properties (of R_X & C_X)

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* Def.: Let b_{ij} be the i, j th element of an $n \times n$ matrix, is non-negative definite if: (NND)

$$\sum_{i=1}^n \sum_{j=1}^n x_i x_j - b_{ij} \geq 0, \forall \text{ real vector } \underline{x} \in \mathbb{R}^n$$

* Note that for "random processes",
 $R_X \leftarrow$ Autocorrelation Matrix

* For any rand. vec. \underline{X} , its correlation matrix is NND

\Rightarrow Proof: Let \underline{x} be an arbitrary vector in \mathbb{R}^n &

$$Y = \underline{x}^T \underline{X} = \underline{X}^T \underline{x} \leftarrow (Y \text{ is a scalar RV})$$

$(1 \times n)(n \times 1) = 1 \times 1$ scalar

$$\Rightarrow 0 \leq E[Y^2] = E[\underline{x}^T \underline{X} \underline{X}^T \underline{x}] = \underline{x}^T E[\underline{X} \underline{X}^T] \underline{x} = \underline{x}^T R_X \underline{x}$$

$$0 \leq \sum_{i=1}^n \sum_{j=1}^n x_i x_j R_{ij} \therefore R_X \text{ is NND (can show same result for } C_X)$$

Gaussian Random Vectors:

* Def.: A Rand. Vec. \underline{X} is gaussian if: $Z = a_0 + \sum_{j=1}^n a_j X_j$ is a gaussian RV $\forall [a_0, \dots, a_n]^T \in \mathbb{R}^{n+1}$

* Instead of deriving a density $f_X(\underline{x})$ for a gaussian Rand. Vec., first find the characteristic function of \underline{X}

Characteristic Function of a Random Vector \underline{X} (we 2 find pdfs of sums of RVs)

$$\Phi_{\underline{X}}(\underline{\omega}) = E\left[e^{i \sum_{j=1}^n \omega_j X_j}\right], \text{ for } \underline{\omega} = [\omega_1, \dots, \omega_n]^T \in \mathbb{R}^n$$

* If $Z = \sum_{j=1}^n X_j$, then $\Phi_Z = E\left[e^{i \omega \sum_{j=1}^n X_j}\right] = \Phi_X(\omega, \dots, \omega)$ (\because every $\omega_j = \omega$)

* If X_1, \dots, X_n are independent then $\Phi_Z(\omega) = \prod_{j=1}^n \Phi_{X_j}(\omega)$

* If X_1, \dots, X_n are i.i.d., then $\Phi_Z(\omega) = (\Phi_X(\omega))^n$ (where Φ_X is the common char. fn of X_1, \dots, X_n)

Char. Fn of Gaussian Rand. Vec. \underline{X}

$$\Phi_{\underline{X}}(\underline{\Omega}) = e^{i\underline{\Omega}^T \underline{\mu}_X - \frac{1}{2} \underline{\Omega}^T \underline{C}_X \underline{\Omega}} \leftarrow \begin{matrix} \text{Char. fn. in terms of the mean-vector} \\ \underline{\mu}_X \text{ \& covar. fn } \underline{C}_X \text{ (2 moments)} \end{matrix}$$

Proof: Let $Z = \sum_{j=1}^n \omega_j X_j$, for $\underline{\Omega} \in \mathbb{R}^n$

then Z is a Gaussian RV $\because \underline{X}$ is a Gaussian Rand. Vector

$$\Rightarrow \Phi_Z(\omega) = \left(e^{i\omega \mu_Z} \right) \left(e^{-\frac{1}{2} \omega^2 \sigma_Z^2} \right), \text{ where } \mu_Z = E[Z] = \sum_{j=1}^n \omega_j \mu_{X_j} = \underline{\Omega}^T \underline{\mu}_X$$

* Z is a linear combination (affine) of X_1, \dots, X_n , which is Gaussian (since \underline{X} is a Gaussian random vector)

$$\sigma_Z^2 = \text{Var}(Z) = \sum_{j=1}^n \sum_{k=1}^n \sigma_{jk} \omega_j \omega_k = \underline{\Omega}^T \underline{C}_X \underline{\Omega}$$

(if $j=k$) $\sigma_{jk} = E[(X_j - \bar{X}_j)(X_k - \bar{X}_k)]$

$$\Rightarrow \Phi_{\underline{X}}(\underline{\Omega}) = E \left[e^{i \sum_{j=1}^n \omega_j X_j} \right] \leftarrow \text{(Char. fn of a Random Vector)}$$

$$= \Phi_Z(\omega)$$

* Substitute $\mu_Z \wedge \sigma_Z^2$ in $\Phi_Z \Rightarrow \Phi_{\underline{X}}(\underline{\Omega}) = \left(e^{i\underline{\Omega}^T \underline{\mu}_X} \right) \left(e^{-\frac{1}{2} \underline{\Omega}^T \underline{C}_X \underline{\Omega}} \right)$

\Rightarrow Density function of a Gaussian random vector \underline{X} is:

(pdf) $f_{\underline{X}}(\underline{x}) = \frac{1}{\sqrt{(2\pi)^n |\underline{C}_X|}} \exp \left[-\frac{1}{2} (\underline{x} - \underline{\mu}_X)^T \underline{C}_X^{-1} (\underline{x} - \underline{\mu}_X) \right]$

(≥ 0) $\because \underline{C}_X^{-1}$ is NND

* In practice, common to seek the most likely value of \underline{x} for a problem which requires maximizing this density function (via maximizing the natural log of the density function)

\Rightarrow Results in a quadratic in the vector \underline{x} containing a unique minimum (\because NND property of \underline{C}_X^{-1}) (Maximization of the likelihood via minimizing the log of this quadratic)